

Chap 6

Part (A): Variance test

Same gaussian models as

Chap 2 : one-sample } focus on σ^2
Chap 5 : two-samples }

One-sample: $y_1, \dots, y_n \stackrel{\text{iid}}{\sim} N(\mu, \sigma^2)$

$$H_0: \sigma^2 = \sigma_0^2$$

$$H_1: \sigma^2 \neq \sigma_0^2$$

Same simulation as Chap 2, but
different stat:

$$s^2 = \frac{1}{n-1} \sum_{i=1}^n (y_i - \bar{y})^2$$

Recall from Chap 3 that

$$V \equiv \frac{(n-1)s^2}{\sigma^2} \sim \chi^2_{n-1}$$

Easier to use $\checkmark$ with closed form calculations.

Confidence Interval for σ^2 .

- Sim: switch σ^2 & s^2 & calc quantiles.

Pr:

$$\begin{aligned}
 1-\alpha &= P(q_{\alpha/2} \leq \sqrt{} \leq q_{1-\alpha/2}) \\
 &= P(q_{\alpha/2} \leq \frac{(n-1)s^2}{\sigma^2} \leq q_{1-\alpha/2}) \\
 &= P\left(\frac{s^2(n-1)}{q_{1-\alpha/2}} \leq \sigma^2 \leq \frac{(n-1)s^2}{q_{\alpha/2}}\right)
 \end{aligned}$$

switched

Two Samples

$$\text{Chap 5: } Y_1, \dots, Y_{n_Y} \stackrel{\text{iid}}{\sim} N(\mu_Y, \sigma_Y^2) \\ \& X_1, \dots, X_{n_X} \stackrel{\text{iid}}{\sim} N(\mu_X, \sigma_X^2)$$

$$H_0: \sigma_Y^2 = \sigma_X^2 \quad \equiv \quad \sigma_Y^2 / \sigma_X^2 = 1$$

$$H_1: \sigma_Y^2 \neq \sigma_X^2$$

$$\text{est } \sigma_Y^2 \text{ w/ } s_Y^2 \quad \& \quad \sigma_X^2 \text{ w/ } s_X^2$$

$$\text{ratio: } f \equiv s_Y^2 / s_X^2 \text{ gives a } \underline{\text{stat.}}$$

$$\text{Chap 3} \quad \left. \begin{aligned} s_Y^2 &\sim \frac{\chi_{n_Y-1}^2 \sigma_Y^2}{n_Y-1} \\ s_X^2 &\sim \frac{\chi_{n_X-1}^2 \sigma_X^2}{n_X-1} \end{aligned} \right) \quad \& \quad \begin{matrix} H_0 \\ \sigma_Y^2 = \sigma_X^2 \end{matrix}$$

$$\text{so } \frac{s_Y^2}{s_X^2} \sim \frac{\chi_{n_Y-1}^2 / (n_Y-1)}{\chi_{n_X-1}^2 / (n_X-1)} \equiv F \stackrel{\text{shown}}{\sim} F_{\substack{n_Y-1, n_X-1 \\ \checkmark \\ \text{DoF}}}$$

Chapter 6

(B) ANOVA: Analysis of ~~Variance~~ Means

Extension of Chap 5 to 3+ (m) samples

$$\text{Model: } Y_{ij} \stackrel{\text{iid}}{\sim} N(\mu_j, \sigma^2)$$

$\uparrow$
pooled

$$i = 1, \dots, n_j$$

$$j = 1, \dots, m$$

$$H_0: \mu_1 = \dots = \mu_m$$

$$H_1: \mu_j \neq \mu_k \quad j \neq k$$

Under H_0 :

$$\bar{y} = \frac{1}{n} \sum_{j=1}^m \sum_{i=1}^{n_j} y_{ij} \equiv \hat{\mu}$$

$$s^2 = \frac{1}{n-1} \sum_{j=1}^m \sum_{i=1}^{n_j} (y_{ij} - \bar{y})^2$$

Unrestricted (under H_1)

Need separate estimates of parameters

$$\hat{\mu}_j \equiv \bar{y}_{.j} = \frac{1}{n_j} \sum_{i=1}^{n_j} y_{ij} \quad j=1, \dots, m$$

$$s_j^2 = \frac{1}{n_j - 1} \sum_{i=1}^{n_j} (y_{ij} - \bar{y}_{.j})^2$$

$$s_{\text{pool}}^2 = \frac{\sum_j (n_j - 1) s_j^2}{n - m} \quad \leftarrow m \equiv \# \text{ of } \bar{y}_{.j}$$

Developing a good statistic for testing involves relating H_0 & H_1 mean-based quantities via their residual sum of squares (SS).

$$\begin{aligned} \text{Hw: } \sum_{j=1}^m \sum_{i=1}^{n_j} (y_{ij} - \bar{y})^2 &= \sum_{j=1}^m \sum_{i=1}^{n_j} (\bar{y}_{.j} - \bar{y})^2 \leftarrow \\ &+ \sum_{j=1}^m \sum_{i=1}^{n_j} (y_{ij} - \bar{y}_{.j})^2 \end{aligned}$$

$\begin{matrix} \text{SST} \rightarrow & \text{SSR} \rightarrow & \text{SSE} \rightarrow \end{matrix}$

$$\text{SST} = \text{SSR} + \text{SSE} \quad \leftarrow$$

Any part over the whole is a
proportion.

proportion.

$$0 < r^2 \equiv \frac{SSR}{SST} = 1 - \frac{SSE}{SST} < 1$$

boiled all my stats into 1 [#].
↑
grand stat.

$$f = \frac{SSR/(m-1)}{SSE/(n-m)} = \frac{MSR}{MSE}$$

under H_0 $F \sim F_{m-1, n-m}$ $\nwarrow$ b/c numer & denom are each χ^2_s